

# On Projections of the Rényi Divergence on Generalized Convex Sets

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**Abstract**—Motivated by a recent result by van Erven and Harremoës, we study a forward projection problem for the Rényi divergence on a particular  $\alpha$ -convex set, termed  $\alpha$ -linear family. The solution to this problem yields a parametric family of probability measures which turns out to be an extension of the exponential family, and it is termed  $\alpha$ -exponential family. An orthogonality relationship between the  $\alpha$ -exponential and  $\alpha$ -linear families is first established and is then used to transform the reverse projection on an  $\alpha$ -exponential family into a forward projection on an  $\alpha$ -linear family. The full paper version of this work is available on the arXiv at <http://arxiv.org/abs/1512.02515>.<sup>1</sup>

**Index Terms** –  $\alpha$ -convex set, relative entropy, variational distance, forward and reverse projections, Rényi divergence, exponential and linear families.

## I. INTRODUCTION

Given a probability measure  $Q$ , and a set of probability measures  $\mathcal{P}$  on an alphabet  $\mathcal{A}$ , a *forward projection* of  $Q$  on  $\mathcal{P}$  is a  $P^* \in \mathcal{P}$  which minimizes the relative entropy  $D(P\|Q)$  subject to  $P \in \mathcal{P}$ . Forward projections appear predominantly in large deviations theory (see, e.g., [4, Chapter 11] in the context of Sanov's theorem and the conditional limit theorem). The forward projection of a generalization of the relative entropy on convex sets has been proposed by Sundaresan in [17] in the context of guessing under source uncertainty, and it was further studied in [12]. In this paper we consider forward projection of Rényi divergence on some generalized convex sets. A physical motivation for such a study stems from a maximum entropy problem proposed by Tsallis in statistical physics [18], [19] (for further details about the connection of this maximum entropy problem to forward projections of the Rényi divergence, the reader is referred to [11]).

The other problem of interest in this paper is the *reverse projection* where the minimization is over the second argument of the divergence measure. This problem is intimately related to the maximum-likelihood estimation and robust statistics. Suppose  $X_1, \dots, X_n$  are i.i.d. samples drawn according to a probability measure which is modeled by a parametric family of probability measures  $\Pi = \{P_\theta : \theta \in \Theta\}$

where  $\Theta$  is a parameter space, and all the members of  $\Pi$  are assumed to have a common finite support  $\mathcal{A}$ . The maximum-likelihood estimator of the given samples (if it exists) is the minimizer of  $D(\hat{P}\|P_\theta)$  subject to  $P_\theta \in \Pi$ , where  $\hat{P}$  is the empirical probability measure of the observed samples [9, Lemma 3.1]. The minimizing probability measure (if exists) is called the reverse projection of  $\hat{P}$  on  $\Pi$ . Other divergences that have natural connection in statistical estimation problems include Hellinger divergence of order  $\frac{1}{2}$  (see, e.g., [3]), Pearson's  $\chi^2$ -divergence, and so on. All these belong to the more general family of Hellinger divergences of order  $\alpha \in (0, \infty)$  (note that these divergences are, up to a positive scaling factor, equal to the power divergences introduced by Cressie and Read [5]); these divergences form a sub-family of  $f$ -divergences which were independently introduced by Ali and Silvey [1] and Csiszár [6]. The Hellinger divergences possess a very good robustness property when a significant fraction of the observed samples are outliers; the textbooks of Basu et al. [2] and Pardo [14] provide a coverage of the developments on the study of inference based on  $f$ -divergences. Since the Rényi divergence is a monotonically increasing function of the Hellinger divergence (see, e.g., [16, (1)]), minimizing a Hellinger divergence of order  $\alpha \in (0, \infty)$  is equivalent to minimizing the Rényi divergence of the same order. This motivates the problem of studying reverse projections of the Rényi divergence in the context of robust statistics.

In a recent work [10, Theorem 14], van Erven and Harremoës proved a Pythagorean property for Rényi divergences of order  $\alpha \in (0, \infty)$  on  $\alpha$ -convex sets. By exploiting this property, we study forward projection of the Rényi divergence on an  $\alpha$ -linear family. The form of forward projection suggests a parametric family of probability measures which turns out to be an extension of the exponential family, and it is termed an  $\alpha$ -exponential family. We show an orthogonality relationship between the  $\alpha$ -linear family and the  $\alpha$ -exponential family. Using this orthogonality property one can transform a reverse projection problem on an  $\alpha$ -exponential family into a forward projection problem on an  $\alpha$ -convex family.

The following is an outline of the paper: Section II provides preliminary material; Section III identifies the form

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of forward projections on  $\alpha$ -linear families, and it proves convergence of an iterated process for forward projections. Section IV proves an orthogonality relationship between the  $\alpha$ -linear and  $\alpha$ -exponential families is established, and this latter property is used to find reverse projections on  $\alpha$ -exponential families. The full paper version of this work is available at [11] which includes additional results and discussions, and all the proofs.

## II. PRELIMINARIES

Unless explicitly mentioned, it is assumed throughout the paper that probability measures are defined on a finite alphabet  $\mathcal{A}$ . Let  $\mathcal{M}$  denote the set of all probability measures on  $\mathcal{A}$ . For  $P \in \mathcal{M}$ , let  $\text{Supp}(P) := \{a \in \mathcal{A} : P(a) > 0\}$ ; for  $\mathcal{P} \subseteq \mathcal{M}$ , let  $\text{Supp}(\mathcal{P})$  be the union of support of members of  $\mathcal{P}$ .

**Definition 1** (Rényi divergence). Let  $\alpha \in (0, 1) \cup (1, \infty)$ . For  $P, Q \in \mathcal{M}$ , the Rényi divergence [15] of order  $\alpha$  from  $P$  to  $Q$  is given by

$$D_\alpha(P\|Q) := \frac{1}{\alpha - 1} \log \left( \sum_a P(a)^\alpha Q(a)^{1-\alpha} \right). \quad (1)$$

If  $\alpha = 1$ , then

$$D_1(P\|Q) := D(P\|Q), \quad (2)$$

which is the analytic extension of  $D_\alpha(P\|Q)$  at  $\alpha = 1$ .

**Definition 2** ( $(\alpha, \lambda)$ -mixture [10]). Let  $P_0, P_1 \in \mathcal{M}$ ,  $\alpha \in (-\infty, 0) \cup (0, \infty)$ , and let  $\lambda \in (0, 1)$ . The  $(\alpha, \lambda)$ -mixture of  $(P_0, P_1)$  is the probability measure  $S_{0,1}$  defined by

$$S_{0,1}(a) := \frac{1}{Z} \left[ (1 - \lambda) P_0(a)^\alpha + \lambda P_1(a)^\alpha \right]^{\frac{1}{\alpha}}, \quad (3)$$

where  $Z$  is the normalizing constant in (3) such that

$$\sum_a S_{0,1}(a) = 1. \quad (4)$$

Here, for simplicity, we suppress the dependence of  $S_{0,1}$  and  $Z$  on  $\alpha, \lambda$ . Note that  $S_{0,1}$  is well-defined as  $Z$  is always positive and finite.

**Definition 3** ( $\alpha$ -convex set). Let  $\alpha \in (-\infty, 0) \cup (0, \infty)$ . A set of probability measures  $\mathcal{P}$  is said to be  $\alpha$ -convex if, for every  $P_0, P_1 \in \mathcal{P}$  and  $\lambda \in (0, 1)$ , the  $(\alpha, \lambda)$ -mixture  $S_{0,1} \in \mathcal{P}$ .

The following is a specific  $\alpha$ -convex set which is of interest in this paper.

**Definition 4** ( $\alpha$ -linear family). Let  $\alpha \in (-\infty, 0) \cup (0, \infty)$ , and  $f_1, \dots, f_k$  be real-valued functions defined on  $\mathcal{A}$ . The  $\alpha$ -linear family determined by  $f_1, \dots, f_k$  is defined to be the following parametric family of probability measures defined on  $\mathcal{A}$ :

$$\mathcal{L}_\alpha := \left\{ P \in \mathcal{M} : P(a) = \left[ \sum_{i=1}^k \theta_i f_i(a) \right]^{\frac{1}{\alpha}}, \quad \underline{\theta} \in \mathbb{R}^k \right\} \quad (5)$$

For typographical convenience, we have suppressed the dependence of  $\mathcal{L}_\alpha$  in  $f_1, \dots, f_k$ . It is easy to see that  $\mathcal{L}_\alpha$  is an

$\alpha$ -convex set. Without loss of generality, we shall assume that  $f_1, \dots, f_k$ , as  $|\mathcal{A}|$ -dimensional vectors, are mutually orthogonal (otherwise, by the Gram-Schmidt procedure, these vectors can be orthogonalized without affecting the corresponding  $\alpha$ -linear family in (5)). Let  $\mathcal{F}$  be the subspace of  $\mathbb{R}^{|\mathcal{A}|}$  spanned by  $f_1, \dots, f_k$ , and let  $\mathcal{F}^\perp$  denote the orthogonal complement of  $\mathcal{F}$ . Hence, there exist  $f_{k+1}, \dots, f_{|\mathcal{A}|}$  such that  $f_1, \dots, f_{|\mathcal{A}|}$  are mutually orthogonal as  $|\mathcal{A}|$ -dimensional vectors, and  $\mathcal{F}^\perp = \text{Span}\{f_{k+1}, \dots, f_{|\mathcal{A}|}\}$ . Consequently, from (5), for  $\alpha \in (0, 1) \cup (1, \infty)$ ,

$$\mathcal{L}_\alpha = \left\{ P \in \mathcal{M} : \sum_a P(a)^\alpha f_i(a) = 0, \quad k+1 \leq i \leq |\mathcal{A}| \right\}. \quad (6)$$

From (6), it is clear that the set  $\mathcal{L}_\alpha$  is closed. Since it is also bounded, it is compact.

## III. FORWARD PROJECTION ON $\alpha$ -LINEAR FAMILY

Let us first recall the Pythagorean property for a Rényi divergence on an  $\alpha$ -convex set. As in the case of relative entropy [9] and relative  $\alpha$ -entropy [13], the Pythagorean property is crucial in establishing orthogonality properties. In the sequel, we assume that  $Q$  is a probability measure with  $\text{Supp}(Q) = \mathcal{A}$ .

In a recent work, van Erven and Harremoës proved a Pythagorean inequality for Rényi divergences on  $\alpha$ -convex sets under the assumption that the forward projection exists (see [10, Theorem 14]). Continuing this study, a sufficient condition for the existence of forward projection is proved in [11] for probability measures on a general alphabet. The following result extends the existence result in [8] for the forward projection of the relative entropy.

**Theorem 1** (Existence of forward  $D_\alpha$ -projection). Let  $\alpha \in (0, \infty)$ , and let  $Q$  be an arbitrary probability measure defined on a set  $\mathcal{A}$ . Let  $\mathcal{P}$  be an  $\alpha$ -convex set of probability measures defined on  $\mathcal{A}$ , and assume that  $\mathcal{P}$  is closed with respect to the total variation distance. If there exists  $P \in \mathcal{P}$  such that  $D_\alpha(P\|Q) < \infty$ , then there exists a forward  $D_\alpha$ -projection of  $Q$  on  $\mathcal{P}$ .

*Proof:* See [11]. For  $\alpha \in (1, \infty)$ , the proof relies on a new Apollonius theorem for the Hellinger divergence, and for  $\alpha \in (0, 1)$ , the proof relies on the Banach-Alaoglu theorem from functional analysis. ■

**Proposition 1** (The Pythagorean property). Let  $\alpha \in (0, 1) \cup (1, \infty)$ , let  $\mathcal{P} \subseteq \mathcal{M}$  be an  $\alpha$ -convex set, and  $Q \in \mathcal{M}$ .

(a) If  $P^*$  is a forward  $D_\alpha$ -projection of  $Q$  on  $\mathcal{P}$ , then

$$D_\alpha(P\|Q) \geq D_\alpha(P\|P^*) + D_\alpha(P^*\|Q) \quad (7)$$

for all  $P \in \mathcal{P}$ . Furthermore, if  $\alpha > 1$ , then  $\text{Supp}(P^*) = \text{Supp}(\mathcal{P})$ .

(b) Conversely, if (7) is satisfied for some  $P^* \in \mathcal{P}$ , then  $P^*$  is the (unique) forward  $D_\alpha$ -projection of  $Q$  on  $\mathcal{P}$ .

*Proof:* See [11]. ■

**Remark 1.** The Pythagorean property (7) holds for probability measures defined on a general alphabet  $\mathcal{A}$ , as proved in [10, Theorem 14]. The novelty here is in the last assertion of (a), which extends the result for relative entropy in [9, Theorem 3.1], for which  $\mathcal{A}$  needs to be finite.

**Corollary 1.** Let  $\alpha \in (0, \infty)$ . If a forward  $D_\alpha$ -projection on an  $\alpha$ -convex set exists, then it is unique.

*Proof:* For  $\alpha = 1$ , since an  $\alpha$ -convex set is particularized to a convex set, this result is known in view of [9, p. 23]. Next, consider the case where  $\alpha \neq 1$ . Let  $P_1^*$  and  $P_2^*$  be forward  $D_\alpha$ -projections of  $Q$  on an  $\alpha$ -convex set  $\mathcal{P}$ . Applying Proposition 1 we have

$$D_\alpha(P_2^* \| Q) \geq D_\alpha(P_2^* \| P_1^*) + D_\alpha(P_1^* \| Q).$$

Since  $D_\alpha(P_1^* \| Q) = D_\alpha(P_2^* \| Q)$ , we must have  $D_\alpha(P_2^* \| P_1^*) = 0$  which implies that  $P_1^* = P_2^*$ . ■

The last assertion in Proposition 1a) states that  $\text{Supp}(P^*) = \text{Supp}(\mathcal{P})$  if  $\alpha \in (1, \infty)$ . The following counterexample, taken from [11], illustrates that this equality does not necessarily hold for  $\alpha \in (0, 1)$ .

**Example 1.** Let  $\mathcal{A} = \{1, 2, 3, 4\}$ ,  $\alpha = \frac{1}{2}$ ,  $f: \mathcal{A} \rightarrow \mathbb{R}$  be given by

$$f(1) = 1, f(2) = -3, f(3) = -5, f(4) = -6 \quad (8)$$

and let  $Q(a) = \frac{1}{4}$  for all  $a \in \mathcal{A}$ . Consider the following  $\alpha$ -linear family:

$$\mathcal{P} := \left\{ P \in \mathcal{M}: \sum_a P(a)^\alpha f(a) = 0 \right\}. \quad (9)$$

Let

$$P^*(1) = \frac{9}{10}, P^*(2) = \frac{1}{10}, P^*(3) = 0, P^*(4) = 0. \quad (10)$$

It is easy to check that  $P^* \in \mathcal{P}$ . Furthermore, setting  $\theta^* = \frac{1}{5}$  and  $Z = \frac{2}{5}$  yields that for  $a \in \{1, 2, 3\}$

$$P^*(a)^{1-\alpha} = Z^{\alpha-1} \left[ Q(a)^{1-\alpha} + (1-\alpha) f(a) \theta^* \right], \quad (11)$$

and

$$P^*(4)^{1-\alpha} > Z^{\alpha-1} \left[ Q(4)^{1-\alpha} + (1-\alpha) f(4) \theta^* \right]. \quad (12)$$

From (9), (11) and (12), it follows that for every  $P \in \mathcal{P}$

$$\sum_{a \in \mathcal{A}} P(a)^\alpha P^*(a)^{1-\alpha} \geq Z^{\alpha-1} \sum_{a \in \mathcal{A}} P(a)^\alpha Q(a)^{1-\alpha}. \quad (13)$$

Furthermore, it can be also verified that

$$Z^{\alpha-1} \sum_{a \in \mathcal{A}} P^*(a)^\alpha Q(a)^{1-\alpha} = 1. \quad (14)$$

Assembling (13) and (14) yields

$$\sum_{a \in \mathcal{A}} P(a)^\alpha P^*(a)^{1-\alpha} \geq \frac{\sum_{a \in \mathcal{A}} P(a)^\alpha Q(a)^{1-\alpha}}{\sum_{a \in \mathcal{A}} P^*(a)^\alpha Q(a)^{1-\alpha}}, \quad (15)$$

which is equivalent to (7). Hence, Proposition 1b) implies that  $P^*$  is the forward  $D_\alpha$ -projection of  $Q$  on  $\mathcal{P}$ . Note,

however, that  $\text{Supp}(P^*) \neq \text{Supp}(\mathcal{P})$ ; to this end, from (9), it can be verified numerically that

$$P = (0.984688, 0.00568298, 0.0041797, 0.00544902) \in \mathcal{P} \quad (16)$$

which implies that  $\text{Supp}(P^*) = \{1, 2\} \subset \mathcal{A}$  whereas  $\text{Supp}(\mathcal{P}) = \mathcal{A}$ .

We shall now focus our attention on forward  $D_\alpha$ -projections on  $\alpha$ -linear families.

**Theorem 2** (Pythagorean equality). Let  $\alpha > 1$ , and let  $P^*$  be the forward  $D_\alpha$ -projection of  $Q$  on an  $\alpha$ -linear family  $\mathcal{L}_\alpha$ . Then,  $P^*$  satisfies (7) with equality, i.e.,

$$D_\alpha(P \| Q) = D_\alpha(P \| P^*) + D_\alpha(P^* \| Q), \quad \forall P \in \mathcal{L}_\alpha. \quad (17)$$

*Proof:* See [11]. ■

In [8, Theorem 3.2], Csiszár proposed a convergent iterative process for finding the forward projection for the relative entropy on a finite intersection of linear families. This result is generalized in this work for the Rényi divergence of order  $\alpha \in (0, \infty)$  on a finite intersection of  $\alpha$ -linear families.

**Theorem 3** (Iterative projections). Let  $\alpha \in (1, \infty)$ . Suppose that  $\mathcal{L}_\alpha^{(1)}, \dots, \mathcal{L}_\alpha^{(m)}$  are  $\alpha$ -linear families, and let

$$\mathcal{P} := \bigcap_{n=1}^m \mathcal{L}_\alpha^{(n)} \quad (18)$$

where  $\mathcal{P}$  is assumed to be a non-empty set. Let  $P_0 = Q$ , and let  $P_n$  for  $n \in \mathbb{N}$  be the forward  $D_\alpha$ -projection of  $P_{n-1}$  on  $\mathcal{L}_\alpha^{(i_n)}$  with  $i_n = n \bmod (m)$ . Then,  $P_n \rightarrow P^*$  (a pointwise convergence) as we let  $n \rightarrow \infty$ .

*Proof:* See [11]. ■

We next identify the form of the forward  $D_\alpha$ -projection on an  $\alpha$ -linear family.

**Theorem 4** (Forward projection on an  $\alpha$ -linear family). Let  $\alpha \in (0, 1) \cup (1, \infty)$ , and let  $P^*$  be the forward  $D_\alpha$ -projection of  $Q$  on an  $\alpha$ -linear family  $\mathcal{L}_\alpha$ . Suppose that

$$\text{Supp}(P^*) = \text{Supp}(\mathcal{L}_\alpha) = \mathcal{A}. \quad (19)$$

Then,

- (a)  $P^*$  satisfies (17).
- (b) There exists  $\theta^* = (\theta_{k+1}^*, \dots, \theta_{|\mathcal{A}|}^*) \in \mathbb{R}^{|\mathcal{A}|-k}$  such that, for all  $a \in \mathcal{A}$ ,

$$\begin{aligned} P^*(a) &= Z(\theta^*)^{-1} \left[ Q(a)^{1-\alpha} + (1-\alpha) \sum_{i=k+1}^{|\mathcal{A}|} \theta_i^* f_i(a) \right]^{\frac{1}{1-\alpha}} \end{aligned} \quad (20)$$

where  $Z(\theta^*)$  is a normalizing constant in (20).

*Proof:* See [11]. ■

**Remark 2.** In view of Example 1, the assumption in (19) does not necessarily hold for  $\alpha \in (0, 1)$ . However, due to

Proposition 1, this assumption necessarily holds for every  $\alpha \in (1, \infty)$ .

For  $\alpha \in (0, \infty)$ , the forward  $D_\alpha$ -projection on an  $\alpha$ -linear family  $\mathcal{L}_\alpha$  motivates the definition of the following parametric family of probability measures. Let  $Q \in \mathcal{M}$ , and let

$$\begin{aligned} \mathcal{E}_\alpha &= \left\{ P \in \mathcal{M} : \right. \\ P(a) &= \frac{1}{Z(\theta)} \left[ Q(a)^{1-\alpha} + (1-\alpha) \sum_{i=k+1}^{|\mathcal{A}|} \theta_i f_i(a) \right]^{\frac{1}{1-\alpha}}, \\ \theta &= (\theta_{k+1}, \dots, \theta_{|\mathcal{A}|}) \in \mathbb{R}^{|\mathcal{A}|-k} \left. \right\}. \end{aligned} \quad (21)$$

It is easy to see that  $\mathcal{E}_\alpha$  is an  $(1-\alpha)$ -convex set. Also, the family  $\mathcal{E}_\alpha$  generalizes the *exponential family* in [9, p. 24]:

$$\begin{aligned} \mathcal{E} &= \left\{ P \in \mathcal{M} : P(a) = Z(\theta)^{-1} Q(a) \exp \left( \sum_{i=k+1}^{|\mathcal{A}|} \theta_i f_i(a) \right), \right. \\ &\quad \left. \theta = (\theta_{k+1}, \dots, \theta_{|\mathcal{A}|}) \in \mathbb{R}^{|\mathcal{A}|-k} \right\}. \end{aligned} \quad (22)$$

This extension is demonstrated in [11]. We shall call the family  $\mathcal{E}_\alpha$  an  $\alpha$ -*exponential family*.<sup>2</sup> It is easy to see that the reference measure  $Q$  in the definition of  $\mathcal{E}_\alpha$  is always a member of  $\mathcal{E}_\alpha$ . As in the case of the exponential family, the  $\alpha$ -exponential family  $\mathcal{E}_\alpha$  also depends on the reference measure  $Q$  only in a loose manner. Any other member of  $\mathcal{E}_\alpha$  could very well play the role of  $Q$  in defining the family. The proof is very similar to the one for the  $\alpha$ -power-law family in [13, Proposition 22]. It should also be noted that, for  $\alpha \in (1, \infty)$ , all members of  $\mathcal{E}_\alpha$  have the same support (i.e., same as the support of  $Q$ ).

#### IV. ORTHOGONALITY OF $\alpha$ -LINEAR AND $\alpha$ -EXPONENTIAL FAMILIES

We first make precise the notion of orthogonality between two sets of probability measures with respect to  $D_\alpha$  ( $\alpha > 0$ ).

**Definition 5** (Orthogonal sets of probability measures). Let  $\alpha \in (0, 1) \cup (1, \infty)$ , and let  $\mathcal{P}$  and  $\mathcal{Q}$  be sets of probability measures. We say that  $\mathcal{P}$  is  $\alpha$ -*orthogonal to*  $\mathcal{Q}$  at  $P^*$  if the following hold:

- (i)  $\mathcal{P} \cap \mathcal{Q} = \{P^*\}$
- (ii)  $D_\alpha(P||Q) = D_\alpha(P||P^*) + D_\alpha(P^*||Q)$  for every  $P \in \mathcal{P}$  and  $Q \in \mathcal{Q}$ .

Note that, when  $\alpha = 1$ , this refers to the orthogonality between the linear and exponential families, which is essentially [9, Corollary 3.1].

We are now ready to state the second main result in [11] namely, the orthogonality between  $\mathcal{L}_\alpha$  and  $\mathcal{E}_\alpha$ .

<sup>2</sup>We emphasize that the  $\alpha$ -*power-law family* proposed in [13, Definition 8] is different extension of the exponential family  $\mathcal{E}$ .

**Theorem 5** (Orthogonality of  $\mathcal{L}_\alpha$  and  $\mathcal{E}_\alpha$ ). Let  $\alpha \in (1, \infty)$ , let  $\mathcal{L}_\alpha$  and  $\mathcal{E}_\alpha$  be given by (5) and (21), respectively, and let  $P^*$  be the forward  $D_\alpha$ -projection of  $Q$  on  $\mathcal{L}_\alpha$ . Then,

- (a)  $\mathcal{L}_\alpha$  is  $\alpha$ -orthogonal to  $\text{cl}(\mathcal{E}_\alpha)$  at  $P^*$ .
- (b) If  $\text{Supp}(\mathcal{L}_\alpha) = \mathcal{A}$ , then  $\mathcal{L}_\alpha$  is  $\alpha$ -orthogonal to  $\mathcal{E}_\alpha$  at  $P^*$ .

*Proof:* See [11]. ■

**Remark 3.** In view of Example 1, if  $\alpha \in (0, 1)$ , then  $\text{Supp}(P^*)$  is not necessarily equal to  $\text{Supp}(\mathcal{L}_\alpha)$ ; this is consistent with Theorem 5 which is stated only for  $\alpha \in (1, \infty)$ . Nevertheless, in view of the proof of Theorem 2, the following holds for  $\alpha \in (0, 1)$ : if  $\text{Supp}(P^*) = \text{Supp}(\mathcal{L}_\alpha) = \mathcal{A}$ , then  $\mathcal{L}_\alpha$  is  $\alpha$ -orthogonal to  $\mathcal{E}_\alpha$  at  $P^*$ .

In [11], Theorem 5 and Remark 3 are applied to find a reverse projection on an  $\alpha$ -exponential family. Before we proceed, we now make precise the definition of a reverse  $D_\alpha$ -projection.

**Definition 6** (Reverse  $D_\alpha$ -projection). Let  $P \in \mathcal{M}$ ,  $\mathcal{Q} \subseteq \mathcal{M}$ , and  $\alpha > 0$ . If there exists  $Q^* \in \mathcal{Q}$  which attains the global minimum of  $D_\alpha(P||Q)$  over all  $Q \in \mathcal{Q}$  and  $D_\alpha(P||Q^*) < \infty$ , then  $Q^*$  is said to be a reverse  $D_\alpha$ -projection of  $P$  on  $\mathcal{Q}$ .

**Theorem 6.** Let  $\alpha \in (0, 1) \cup (1, \infty)$ , and let  $\mathcal{E}_\alpha$  be an  $\alpha$ -exponential family determined by  $Q, f_{k+1}, \dots, f_{|\mathcal{A}|}$ . Let  $X_1, \dots, X_n$  be i.i.d. samples drawn at random according to a probability measure in  $\mathcal{E}_\alpha$ . Let  $\hat{P}_n$  be the empirical probability measure of  $X_1, \dots, X_n$ , and let  $P_n^*$  be the forward  $D_\alpha$ -projection of  $Q$  on the  $\alpha$ -linear family

$$\mathcal{L}_\alpha^{(n)} := \left\{ P \in \mathcal{M} : \sum_a P(a) \hat{f}_i^{(n)}(a) = 0, k+1 \leq i \leq |\mathcal{A}| \right\}, \quad (23)$$

where

$$\hat{f}_i^{(n)}(a) := f_i(a) - \hat{\eta}_i^{(n)} Q(a)^{1-\alpha}, \quad \forall a \in \mathcal{A} \quad (24)$$

with

$$\hat{\eta}_i^{(n)} := \frac{\sum_a \hat{P}_n(a) f_i(a)}{\sum_a \hat{P}_n(a)^\alpha Q(a)^{1-\alpha}}, \quad i \in \{k+1, \dots, |\mathcal{A}|\}. \quad (25)$$

Then, the following hold:

- (a) If  $\text{Supp}(P_n^*) = \text{Supp}(\mathcal{L}_\alpha^{(n)}) = \mathcal{A}$ , then  $P_n^*$  is the reverse  $D_\alpha$ -projection of  $\hat{P}_n$  on  $\mathcal{E}_\alpha$ .
- (b) For  $\alpha \in (1, \infty)$ , if  $\text{Supp}(\mathcal{L}_\alpha^{(n)}) \neq \mathcal{A}$ , then the reverse  $D_\alpha$ -projection of  $\hat{P}_n$  on  $\mathcal{E}_\alpha$  does not exist. Nevertheless,  $P_n^*$  is the reverse  $D_\alpha$ -projection of  $\hat{P}_n$  on  $\text{cl}(\mathcal{E}_\alpha)$ .

*Proof:* See [11]. ■

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