

# Arimoto-Rényi Conditional Entropy and Bayesian $M$ -ary Hypothesis Testing

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## Hypothesis Testing

- Bayesian  $M$ -ary hypothesis testing:
  - ▶  $X$  is a random variable taking values on  $\mathcal{X}$  with  $|\mathcal{X}| = M$ ;
  - ▶ a prior distribution  $P_X$  on  $\mathcal{X}$ ;
  - ▶  $M$  hypotheses for the  $\mathcal{Y}$ -valued data  $\{P_{Y|X=m}, m \in \mathcal{X}\}$ .

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- ▶  $M$  hypotheses for the  $\mathcal{Y}$ -valued data  $\{P_{Y|X=m}, m \in \mathcal{X}\}$ .

- $\varepsilon_{X|Y}$ : the minimum probability of error of  $X$  given  $Y$

- ▶ achieved by the *maximum-a-posteriori* (MAP) decision rule. Hence,

$$\varepsilon_{X|Y} = \mathbb{E} \left[ 1 - \max_{x \in \mathcal{X}} P_{X|Y}(x|Y) \right] \quad (1)$$

$$= 1 - \sum_{y \in \mathcal{Y}} \max_{x \in \mathcal{X}} P_{X,Y}(x, y). \quad (2)$$

where (2) holds when  $Y$  is discrete.

## Example

Let  $X$  and  $Y$  be random variables defined on the set  $\mathcal{A} = \{1, 2, 3\}$ , and let

$$[P_{XY}(x, y)]_{(x, y) \in \mathcal{A}^2} = \frac{1}{45} \begin{pmatrix} 8 & 1 & 6 \\ 3 & 5 & 7 \\ 4 & 9 & 2 \end{pmatrix}. \quad (3)$$

Then,

$$\varepsilon_{X|Y} = 1 - \left( \frac{8}{45} + \frac{9}{45} + \frac{7}{45} \right) = \frac{7}{15}. \quad (4)$$

Interplay  $\varepsilon_{X|Y} \longleftrightarrow$  information measures

- Bounds on  $\varepsilon_{X|Y}$  involving information measures exist in the literature. Those works attest that there is a considerable motivation for studying the relationships between  $\varepsilon_{X|Y}$  and information measures.
- $\varepsilon_{X|Y}$  is rarely directly computable, and the best bounds are information theoretic.

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  - ▶ proofs of coding theorems.

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  - ▶ the analysis of  $M$ -ary hypothesis testing
  - ▶ proofs of coding theorems.
- In this talk, we introduce:

upper and lower bounds on  $\varepsilon_{X|Y}$  in terms of the *Arimoto-Rényi* conditional entropy  $H_\alpha(X|Y)$  of any order  $\alpha$ , and apply them in coding.

## The Rényi Entropy

## Definition

Let  $P_X$  be a probability distribution on a discrete set  $\mathcal{X}$ . The Rényi entropy of order  $\alpha \in (0, 1) \cup (1, \infty)$  of  $X$  is defined as

$$H_\alpha(X) = \frac{1}{1-\alpha} \log \sum_{x \in \mathcal{X}} P_X^\alpha(x) \quad (5)$$

By its continuous extension,

$$H_0(X) = \log |\{x \in \mathcal{X} : P_X(x) > 0\}|, \quad (6)$$

$$H_1(X) = H(X), \quad (7)$$

$$H_\infty(X) = \log \frac{1}{p_{\max}} \quad (8)$$

where  $p_{\max}$  is the largest of the masses of  $X$ .

## The Binary Rényi Divergence

### Definition

For  $\alpha \in (0, 1) \cup (1, \infty)$ , the **binary Rényi divergence of order  $\alpha$**  is given by

$$d_\alpha(p\|q) = \frac{1}{\alpha - 1} \log\left(p^\alpha q^{1-\alpha} + (1-p)^\alpha (1-q)^{1-\alpha}\right). \quad (9)$$

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$$\lim_{\alpha \uparrow 1} d_\alpha(p\|q) = d(p\|q) = p \log \frac{p}{q} + (1-p) \log \frac{1-p}{1-q}. \quad (10)$$

## Rényi Conditional Entropy ?

- If we mimic the definition of  $H(X|Y)$  and define conditional Rényi entropy as

$$\sum_{y \in \mathcal{Y}} P_Y(y) H_\alpha(X|Y = y),$$

we find that, for  $\alpha \neq 1$ , the conditional version may be larger than  $H_\alpha(X)$  !

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- To remedy this situation, Arimoto introduced a notion of conditional Rényi entropy,  $H_\alpha(X|Y)$  (named **Arimoto-Rényi conditional entropy**), which is upper bounded by  $H_\alpha(X)$ .

## The Arimoto-Rényi Conditional Entropy (cont.)

## Definition

Let  $P_{XY}$  be defined on  $\mathcal{X} \times \mathcal{Y}$ , where  $X$  is a discrete random variable.

- If  $\alpha \in (-\infty, 0) \cup (0, 1) \cup (1, \infty)$ , then

$$H_\alpha(X|Y) = \frac{\alpha}{1-\alpha} \log \mathbb{E} \left[ \left( \sum_{x \in \mathcal{X}} P_{X|Y}^\alpha(x|Y) \right)^{\frac{1}{\alpha}} \right] \quad (11)$$

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$$= \frac{\alpha}{1-\alpha} \log \sum_{y \in \mathcal{Y}} P_Y(y) \exp \left( \frac{1-\alpha}{\alpha} H_\alpha(X|Y=y) \right), \quad (12)$$

where (12) applies if  $Y$  is a discrete random variable.

## The Arimoto-Rényi Conditional Entropy (cont.)

- By its continuous extension,

$$H_0(X|Y) = \text{ess sup } H_0(P_{X|Y}(\cdot|Y)) \quad (13)$$

$$= \max_{y \in \mathcal{Y}} H_0(X | Y = y), \quad (14)$$

$$H_1(X|Y) = H(X|Y), \quad (15)$$

$$H_\infty(X|Y) = \log \frac{1}{\mathbb{E} \left[ \max_{x \in \mathcal{X}} P_{X|Y}(x|Y) \right]} \quad (16)$$

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## Monotonicity Properties

- $H_\alpha(X|Y)$  is monotonically **decreasing** in  $\alpha$  throughout the real line.
- $\frac{\alpha-1}{\alpha} H_\alpha(X|Y)$  is monotonically **increasing** in  $\alpha$  on  $(0, \infty)$  &  $(-\infty, 0)$ .

## Fano's Inequality

Let  $X$  take values in  $|\mathcal{X}| = M$ , then

$$H(X|Y) \leq h(\varepsilon_{X|Y}) + \varepsilon_{X|Y} \log(M - 1) \quad (17)$$

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$$= \log M - d\left(\varepsilon_{X|Y} \parallel 1 - \frac{1}{M}\right) \quad (18)$$

- (18) is not nearly as popular as (17);
- (18) turns out to be the version that admits an elegant (although not immediate) generalization to the Arimoto-Rényi conditional entropy.

## Generalization of Fano's Inequality

- It is easy to get Fano's inequality by averaging  $H(X|Y = y)$  with respect to the observation  $y$ :

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- This simple route is not viable in the case of  $H_\alpha(X|Y)$  since it is not an average of Rényi entropies of conditional distributions:

$$H_\alpha(X|Y) \neq \sum_{y \in \mathcal{Y}} P_Y(y) H_\alpha(X|Y = y), \quad \alpha \neq 1. \quad (19)$$

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- The standard proof of Fano's inequality, also fails for  $H_\alpha(X|Y)$  of order  $\alpha \neq 1$  since it **does not satisfy the chain rule**.

## Generalization of Fano's Inequality (cont.)

Before we generalize Fano's inequality by linking  $\varepsilon_{X|Y}$  with  $H_\alpha(X|Y)$  for  $\alpha \in [0, \infty)$ , note that for  $\alpha = \infty$ , the following equality holds:

$$\varepsilon_{X|Y} = 1 - \exp(-H_\infty(X|Y)). \quad (20)$$

## Generalization of Fano's Inequality (cont.)

## Lemma

Let  $\alpha \in (0, 1) \cup (1, \infty)$  and  $(\beta, \gamma) \in (0, \infty)^2$ . Then,

$$f_{\alpha, \beta, \gamma}(u) = (\gamma(1-u)^\alpha + \beta u^\alpha)^{\frac{1}{\alpha}}, \quad u \in [0, 1] \quad (21)$$

is

- strictly convex for  $\alpha \in (1, \infty)$ ;
- strictly concave for  $\alpha \in (0, 1)$ .

$$f''_{\alpha, \beta, \gamma}(u) = (\alpha - 1)\beta\gamma \left( \gamma(1-u)^\alpha + \beta u^\alpha \right)^{\frac{1}{\alpha} - 2} (u(1-u))^{\alpha - 2} \quad (22)$$

which is strictly negative if  $\alpha \in (0, 1)$ , and strictly positive if  $\alpha \in (1, \infty)$ .

## Generalization of Fano's Inequality (cont.)

## Theorem

Let  $P_{XY}$  be a probability measure defined on  $\mathcal{X} \times \mathcal{Y}$  with  $|\mathcal{X}| = M < \infty$ . For all  $\alpha \in (0, \infty)$ ,

$$H_\alpha(X|Y) \leq \log M - d_\alpha(\varepsilon_{X|Y} \| 1 - \frac{1}{M}). \quad (23)$$

Equality holds in (23) if and only if, for all  $y$ ,

$$P_{X|Y}(x|y) = \begin{cases} \frac{\varepsilon_{X|Y}}{M-1}, & x \neq \mathcal{L}^*(y) \\ 1 - \varepsilon_{X|Y}, & x = \mathcal{L}^*(y) \end{cases} \quad (24)$$

where  $\mathcal{L}^* : \mathcal{Y} \rightarrow \mathcal{X}$  is a deterministic MAP decision rule.

## Generalization of Fano's Inequality (cont.)

If  $X, Y$  are vectors of dimension  $n$ , then  $\varepsilon_{X|Y} \rightarrow 0 \Rightarrow \frac{1}{n}H(X|Y) \rightarrow 0$ .  
However, the picture with  $H_\alpha(X|Y)$  is more nuanced !

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### Theorem

#### Assume

- $\{X_n\}$  is a sequence of random variables;
- $X_n$  takes values on  $\mathcal{X}_n$  such that  $|\mathcal{X}_n| \leq M^n$  for  $M \geq 2$  and all  $n$ ;
- $\{Y_n\}$  is a sequence of random variables, for which  $\varepsilon_{X_n|Y_n} \rightarrow 0$ .

- a) If  $\alpha \in (1, \infty]$ , then  $H_\alpha(X_n|Y_n) \rightarrow 0$ ;
- b) If  $\alpha = 1$ , then  $\frac{1}{n}H(X_n|Y_n) \rightarrow 0$ ;
- c) If  $\alpha \in [0, 1)$ , then  $\frac{1}{n}H_\alpha(X_n|Y_n)$  is upper bounded by  $\log M$ ;  
 nevertheless, it does not necessarily tend to 0.

Lower Bound on  $H_\alpha(X|Y)$ 

## Theorem

If  $\alpha \in (0, 1) \cup (1, \infty)$ , then

$$\frac{\alpha}{1-\alpha} \log g_\alpha(\varepsilon_{X|Y}) \leq H_\alpha(X|Y), \quad (25)$$

with the piecewise linear function

$$g_\alpha(t) = \left( k(k+1)^{\frac{1}{\alpha}} - k^{\frac{1}{\alpha}}(k+1) \right) t + k^{\frac{1}{\alpha}+1} - (k-1)(k+1)^{\frac{1}{\alpha}} \quad (26)$$

on the interval  $t \in \left[ 1 - \frac{1}{k}, 1 - \frac{1}{k+1} \right)$  for  $k \in \{1, 2, \dots\}$ .

- Not restricted to finite  $M$ .

## Proof Outline

## Lemma

Let  $X$  be a discrete random variable attaining maximal mass  $p_{\max}$ . Then, for  $\alpha \in (0, 1) \cup (1, \infty)$ ,

$$H_\alpha(X) \geq s_\alpha(\varepsilon_X) \quad (27)$$

where  $\varepsilon_X = 1 - p_{\max}$  is the minimum error probability of guessing  $X$ , and  $s_\alpha: [0, 1) \rightarrow [0, \infty)$  is given by

$$s_\alpha(x) := \frac{1}{1-\alpha} \log \left( \left\lfloor \frac{1}{1-x} \right\rfloor (1-x)^\alpha + \left( 1 - (1-x) \left\lfloor \frac{1}{1-x} \right\rfloor \right)^\alpha \right).$$

Equality holds in (27) if and only if  $P_X$  has  $\left\lfloor \frac{1}{p_{\max}} \right\rfloor$  masses equal to  $p_{\max}$ .

The proof relies on the Schur-concavity of  $H_\alpha(\cdot)$ .

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For every  $y \in \mathcal{Y}$ , the lemma yields  $H_\alpha(X | Y = y) \geq s_\alpha(\varepsilon_{X|Y}(y))$ .

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For  $\alpha \in (0, 1)$ , let  $f_\alpha: [0, 1) \rightarrow [1, \infty)$  be defined as

$$f_\alpha(x) = \exp\left(\frac{1-\alpha}{\alpha} s_\alpha(x)\right)$$

- $g_\alpha$  is the piecewise linear function which coincides with  $f_\alpha$  at all points  $1 - \frac{1}{k}$  for  $k \in \mathbb{N}$ ;
- $g_\alpha$  is the **lower convex envelope** of  $f_\alpha$ ;

$$\begin{aligned} H_\alpha(X|Y) &\geq \frac{\alpha}{1-\alpha} \log \mathbb{E} [f_\alpha(\varepsilon_{X|Y}(Y))] \quad (\text{Lemma; } f_\alpha \text{ increasing}) \\ &\geq \frac{\alpha}{1-\alpha} \log \mathbb{E} [g_\alpha(\varepsilon_{X|Y}(Y))] \quad (g_\alpha \leq f_\alpha) \\ &\geq \frac{\alpha}{1-\alpha} \log g_\alpha(\varepsilon_{X|Y}) \quad (\text{Jensen}) \end{aligned}$$

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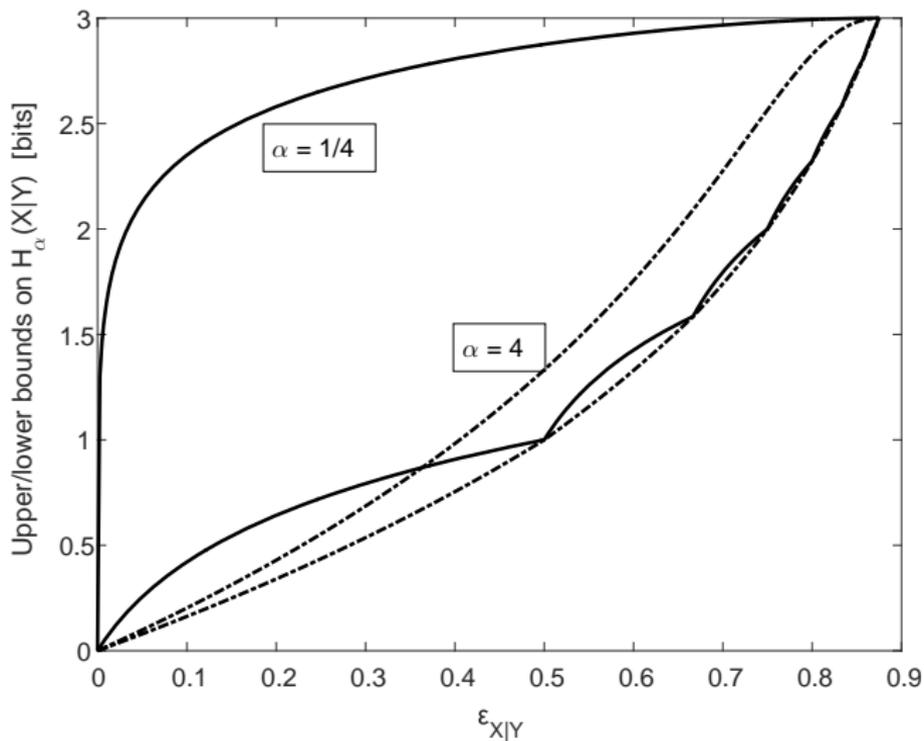
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For  $\alpha \in (1, \infty)$ ,  $-g_\alpha$  is the lower convex envelope of  $-f_\alpha$ , and  $f_\alpha$  is monotonically decreasing. Proof is similar.

$$H_\alpha(X|Y) \longleftrightarrow \varepsilon_{X|Y}$$



## Asymptotic Tightness

Both upper and lower bounds on  $\varepsilon_{X|Y}$  are asymptotically tight as  $\alpha \rightarrow \infty$ .

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## Special cases

As  $\alpha \rightarrow 1$ , we get existing bounds as special cases:

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- Its counterpart by Kovalevsky ('68), and Tebbe and Dwyer ('68).

## Upper bound on $\varepsilon_{X|Y}$

The most useful domain of applicability of the counterpart to the generalization of Fano's inequality is  $\varepsilon_{X|Y} \in [0, \frac{1}{2}]$ , in which case the lower bound specializes to ( $k = 1$ )

$$\frac{\alpha}{1-\alpha} \log\left(1 + \left(2^{\frac{1}{\alpha}} - 2\right)\varepsilon_{X|Y}\right) \leq H_{\alpha}(X|Y). \quad (28)$$

## List Decoding

- Decision rule outputs a list of choices.
- The extension of Fano's inequality to list decoding, expressed in terms of the conditional Shannon entropy, was initiated by Ahlswede, Gacs and Körner ('66).
- Useful for proving converse results.

## Generalization of Fano's Inequality for List Decoding

- A generalization of Fano's inequality for list decoding of size  $L$  is

$$H(X|Y) \leq \log M - d(P_{\mathcal{L}} \| 1 - \frac{L}{M}), \quad (29)$$

where  $P_{\mathcal{L}}$  denotes the probability of  $X$  not being in the list.

- Averaging a conditional version of  $H_{\alpha}(X|Y = y)$  with respect to the observation is not viable in the case of  $H_{\alpha}(X|Y)$  with  $\alpha \neq 1$ .

## Generalization of Fano's Inequality for List Decoding (cont.)

## Theorem (Fixed List Size)

Let  $P_{XY}$  be a probability measure defined on  $\mathcal{X} \times \mathcal{Y}$  where  $|\mathcal{X}| = M$ . Consider a decision rule<sup>a</sup>  $\mathcal{L}: \mathcal{Y} \rightarrow \binom{\mathcal{X}}{L}$ , and denote the decoding error probability by  $P_{\mathcal{L}} = \mathbb{P}[X \notin \mathcal{L}(Y)]$ . Then, for all  $\alpha \in (0, 1) \cup (1, \infty)$ ,

$$H_{\alpha}(X|Y) \leq \log M - d_{\alpha}(P_{\mathcal{L}} \| 1 - \frac{L}{M}) \quad (30)$$

with equality in (30) if and only if

$$P_{X|Y}(x|y) = \begin{cases} \frac{P_{\mathcal{L}}}{M-L}, & x \notin \mathcal{L}(y) \\ \frac{1-P_{\mathcal{L}}}{L}, & x \in \mathcal{L}(y). \end{cases} \quad (31)$$

<sup>a</sup> $\binom{\mathcal{X}}{L}$  stands for the set of all subsets of  $\mathcal{X}$  with cardinality  $L$ , with  $L \leq |\mathcal{X}|$ .

## Arimoto-Rényi Conditional Entropy Averaged over Codebook Ensembles

- Consider the channel coding setup with a code ensemble  $\mathcal{C}$ , over which we are interested in averaging the Arimoto-Rényi conditional entropy of the channel input given the channel output.
- Denote such averaged quantity by

$$\mathbb{E}_{\mathcal{C}}[H_{\alpha}(X^n|Y^n)]$$

where  $X^n = (X_1, \dots, X_n)$  and  $Y^n = (Y_1, \dots, Y_n)$ .

- Some motivation for this study:
  - ▶ The normalized equivocation  $\frac{1}{n}H(X^n|Y^n)$  was used by Shannon to prove that reliable communication is impossible at rates above capacity;
  - ▶ The asymptotic convergence to zero of the equivocation  $H(X^n|Y^n)$  at rates below capacity was studied by Feinstein.

## Coding Theorem 1 (Feder and Merhav, 1994)

For a DMC with transition probability matrix  $P_{Y|X}$ , the conditional entropy of the transmitted codeword given the channel output, averaged over a random coding selection with per-letter distribution  $P_X$  such that  $I(P_X, P_{Y|X}) > 0$ , is bounded (in nats) by

$$\mathbb{E}_{\mathcal{C}} [H(X^n|Y^n)] \leq \left(1 + \frac{1}{\rho^*(R, P_X)}\right) \exp(-nE_r(R, P_X))$$

with

- $R = \frac{\log M}{n} \leq I(P_X, P_{Y|X})$ ;
- $E_r$  is the **random-coding error exponent**, given by

$$E_r(R, P_X) = \max_{\rho \in [0,1]} \rho \left( I_{\frac{1}{1+\rho}}(P_X, P_{Y|X}) - R \right); \quad (32)$$

- the argument that maximizes (32) is denoted by  $\rho^*(R, P_X)$ .

## Coding Theorem 2 (ISSV, 2017)

The following results hold under the setting in the previous theorem:

- For all  $\alpha > 0$ , and rates  $R$  below the channel capacity  $C$ ,

$$\limsup_{n \rightarrow \infty} -\frac{1}{n} \log \mathbb{E}_{\mathcal{C}} [H_{\alpha}(X^n | Y^n)] \leq E_{\text{sp}}(R), \quad (33)$$

where  $E_{\text{sp}}(\cdot)$  denotes **the sphere-packing error exponent**

$$E_{\text{sp}}(R) = \sup_{\rho \geq 0} \rho \left( \max_{Q_X} I_{\frac{1}{1+\rho}}(Q_X, P_{Y|X}) - R \right) \quad (34)$$

with the maximization in the right side of (34) over all single-letter distributions  $Q_X$  defined on the input alphabet.

## Coding Theorem 2 (ISSV '17, cont.)

- For all  $\alpha \in (0, 1)$ ,

$$\liminf_{n \rightarrow \infty} -\frac{1}{n} \log \mathbb{E}_{\mathcal{C}} [H_{\alpha}(X^n | Y^n)] \geq \alpha E_r(R, P_X) - (1 - \alpha)R, \quad (35)$$

provided that

$$R < R_{\alpha}(P_X, P_{Y|X}) \quad (36)$$

where  $R_{\alpha}(P_X, P_{Y|X})$  is the unique solution  $r \in (0, I(P_X, P_{Y|X}))$  to

$$E_r(r, P_X) = \left( \frac{1}{\alpha} - 1 \right) r. \quad (37)$$

## Coding Theorem 2 (ISSV '17, cont.)

- The rate  $R_\alpha(P_X, P_{Y|X})$  is monotonically increasing and continuous in  $\alpha \in (0, 1)$ , and

$$\lim_{\alpha \downarrow 0} R_\alpha(P_X, P_{Y|X}) = 0, \quad (38)$$

$$\lim_{\alpha \uparrow 1} R_\alpha(P_X, P_{Y|X}) = I(P_X, P_{Y|X}). \quad (39)$$

## Coding Theorem 3 (ISSV '17, cont.)

Let  $P_{Y|X}$  be the transition probability matrix of a memoryless binary-input output-symmetric channel, and let  $P_X^* = [\frac{1}{2} \ \frac{1}{2}]$ . Let  $R_c$ ,  $R_0$ , and  $C$  denote the critical and cutoff rates and the channel capacity, respectively, and let

$$\alpha_c = \frac{R_c}{R_0} \in (0, 1). \quad (40)$$

The rate  $R_\alpha = R_\alpha(P_X^*, P_{Y|X})$ , with the symmetric input distribution  $P_X^*$ , can be expressed as follows:

- a) for  $\alpha \in (0, \alpha_c]$ ,  $R_\alpha = \alpha R_0$ ;
- b) for  $\alpha \in (\alpha_c, 1)$ ,  $R_\alpha \in (R_c, C)$  is the solution to  $E_{sp}(r) = (\frac{1}{\alpha} - 1) r$ ;
- c)  $R_\alpha$  is continuous, monotonically increasing in  $\alpha \in [\alpha_c, 1)$  from  $R_c$  to  $C$ .

Example: BSC( $\delta$ )

- Consider a BSC with crossover probability  $\delta$ , and let  $P_X = \left[\frac{1}{2} \ \frac{1}{2}\right]$ .
- the cutoff rate, critical rate and capacity (in bits) are given by

$$R_0 = 1 - \log(1 + \sqrt{4\delta(1-\delta)}), \quad (41)$$

$$R_c = 1 - h\left(\frac{\sqrt{\delta}}{\sqrt{\delta} + \sqrt{1-\delta}}\right), \quad (42)$$

$$C = I(P_X, P_{Y|X}) = 1 - h(\delta). \quad (43)$$

- The sphere-packing error exponent is given by

$$E_{\text{sp}}(R) = d(\delta_{\text{GV}}(R) \parallel \delta) \quad (44)$$

where the **normalized Gilbert-Varshamov distance** is denoted by

$$\delta_{\text{GV}}(R) = h^{-1}(1 - R). \quad (45)$$

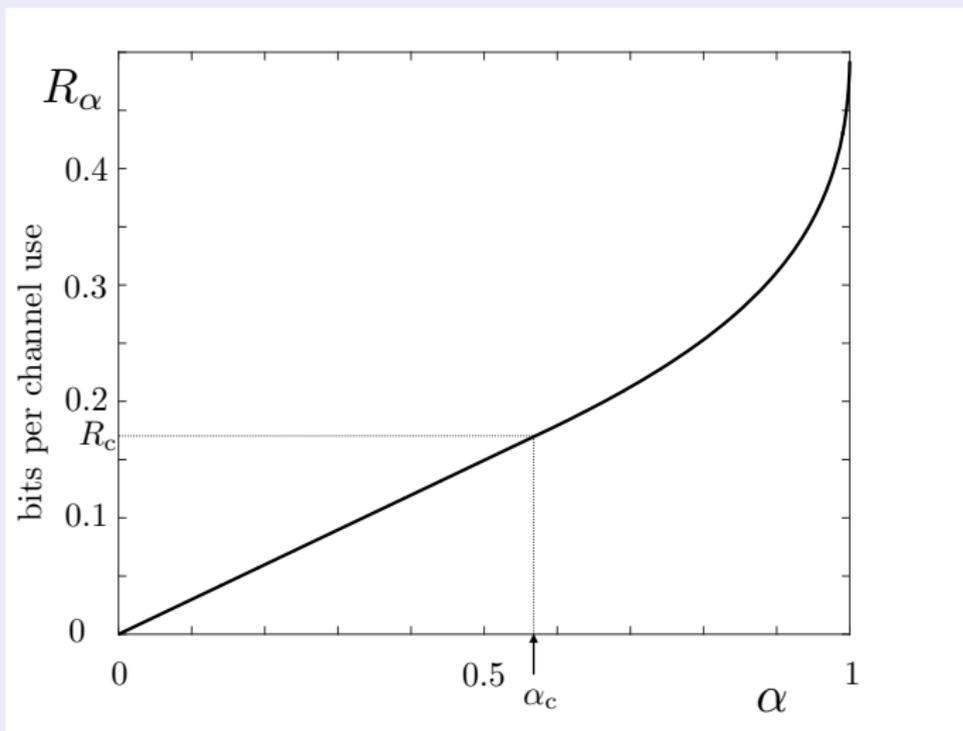
Example: BSC( $\delta$ ) (cont.)

Figure: The rate  $R_\alpha$  for  $\alpha \in (0, 1)$  for BSC( $\delta$ ) with crossover prob.  $\delta = 0.110$ .

## Conclusions

- We have shown new bounds on the **minimum Bayesian error prob.**  $\epsilon_{X|Y}$  of  $M$ -ary hypothesis testing.
- Our major focus has been the **Arimoto-Rényi conditional entropy** of the hypothesis index given the observation.

## Conclusions

- We have shown new bounds on the **minimum Bayesian error prob.**  $\varepsilon_{X|Y}$  of  $M$ -ary hypothesis testing.
- Our major focus has been the **Arimoto-Rényi conditional entropy** of the hypothesis index given the observation.
- Changing the conventional form of Fano's inequality from

$$H(X|Y) \leq h(\varepsilon_{X|Y}) + \varepsilon_{X|Y} \log(M - 1) \quad (46)$$

$$= \log M - d(\varepsilon_{X|Y} \| 1 - \frac{1}{M}) \quad (47)$$

to the right side of (47), where  $d(\cdot \| \cdot)$  is the binary relative entropy, allows a natural generalization where the Arimoto-Rényi conditional entropy of an arbitrary positive order  $\alpha$  is upper bounded by

$$H_\alpha(X|Y) \leq \log M - d_\alpha(\varepsilon_{X|Y} \| 1 - \frac{1}{M}) \quad (48)$$

with  $d_\alpha(\cdot \| \cdot)$  denoting the **binary Rényi divergence**.

## Conclusions (Cont.)

- The Schur-concavity of the Rényi entropy yields a lower bound on  $H_\alpha(X|Y)$  in terms of  $\varepsilon_{X|Y}$ , which holds even if  $M = \infty$ . It recovers existing bounds by letting  $\alpha \rightarrow 1$ .

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- Our techniques were extended to list decoding with a fixed list size, generalizing all the  $H_\alpha(X|Y) - \varepsilon_{X|Y}$  bounds to that setting.
- **Application:** We analyzed the exponentially vanishing decay of the Arimoto-Rényi conditional entropy of the transmitted codeword given the channel output for DMCs and random coding ensembles.

## Further Results in This Work

- Explicit lower bounds on  $\varepsilon_{X|Y}$  as a function of  $H_\alpha(X|Y)$  for an arbitrary  $\alpha$  (also, for  $\alpha < 0$ ).
- Explicit lower bounds on the list decoding error probability for fixed list size as a function of  $H_\alpha(X|Y)$  for an arbitrary  $\alpha$  (also, for  $\alpha < 0$ ).
- We also explored some facets of the role of binary hypothesis testing in analyzing  $M$ -ary Bayesian hypothesis testing problems, and have shown new bounds in terms of Rényi divergence.

## Journal Paper

I. Sason and S. Verdú, “Arimoto-Rényi conditional entropy and Bayesian  $M$ -ary hypothesis testing,” to appear in the *IEEE Trans. on Information Theory*. [Online]. Available at <https://arxiv.org/abs/1701.01974>.